

COSMIC VARIANCE IN THE GREAT OBSERVATORIES ORIGINS DEEP SURVEY

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ABSTRACT

Cosmic variance is the uncertainty in observational estimates of the volume density of extragalactic objects such as galaxies or quasars arising from the underlying large-scale density fluctuations. This is often a significant source of uncertainty, especially in deep galaxy surveys, which tend to cover relatively small areas. We present estimates of the relative cosmic variance for one-point statistics (i.e. number densities) for typical scales and volumes sampled by the Great Observatories Origins Deep Survey (GOODS). We use two approaches: for objects with a known two-point correlation function that is well-approximated by a power law, one can use the standard analytic formalism to calculate the cosmic variance (in excess of shot noise). We use this approach to estimate the cosmic variance for several populations that are being studied in the GOODS program: Extremely Red Objects (ERO) at $z \sim 1$, and Lyman Break Galaxies (LBG) at $z \sim 3$ and $z \sim 4$, using clustering information for similar populations in the literature. For populations with unknown clustering, one can use predictions from Cold Dark Matter theory to obtain a rough estimate of the variance as a function of number density. We present a convenient plot which allows one to use this approach to read off the cosmic variance for a population with a known mean redshift and estimated number density. We conclude that for the volumes sampled by GOODS, cosmic variance is a significant source of uncertainty for strongly clustered objects (~ 40 -60% for EROs) and less serious for less clustered objects, ~ 10 -20% for LBGs.

Subject headings: galaxies: statistics — large scale structure of universe

1. INTRODUCTION

The number density of observed extragalactic populations in the Universe is a fundamental property which may hold clues to the nature of the objects. However, observational estimates of the number density of any clustered population are plagued by uncertainty due to *cosmic variance*, the field-to-field variation (in excess of Poisson shot noise) due to large scale structure. Clearly, if one can sample a volume that is very large compared with the intrinsic clustering scale of the objects in question, cosmic variance will be insignificant. In practice, especially in high-redshift studies, the volumes sampled are small enough that cosmic variance is often a significant source of uncertainty. Perhaps the majority of published cosmological number densities and related quantities (e.g. luminosity functions, integrated luminosity densities, etc.) do not properly account for cosmic variance in their quoted error budgets.

Cosmic variance has frequently been invoked as a motivation for carrying out deep pencil-beam surveys along multiple sightlines. While the term “cosmic variance” is generally understood, the effects are of course dependent on the clustering properties of the sources of interest, a fact that is often lost in discussions of deep survey strategy. With the availability of the GOODS data, it seems appropriate to cast this variance in practical terms, calculating explicitly the expected uncertainties due to clustering for various source populations under study. A simple exposition of this cosmic variance may be useful both to researchers using the GOODS data and to those planning future studies.

The mean $\langle N \rangle$ and variance $\langle N^2 \rangle$ are the first and second moments of the probability distribution function $P_N(V)$, which represents the probability of counting N objects within a volume V . We define the *relative* cosmic variance:

$$\sigma_v^2 \equiv \frac{\langle N^2 \rangle - \langle N \rangle^2}{\langle N \rangle^2} - \frac{1}{\langle N \rangle}. \quad (1)$$

Note that the last term is the usual correction for Poisson shot noise, which for the samples considered here will typically be negligible. In any case, it is relatively straightforward to perform this correction, so we do not discuss this term further. In the general hierarchical scenario of structure formation, in which density perturbations grow via gravitational instability, P_N is expected to have non-zero higher moments (e.g. skewness and kurtosis). For a detailed and general treatment of the cosmic error, see Colombi et al. (2000), Szapudi et al. (1999, 2000), and references therein. Here, we concentrate solely on one-point statistics (i.e. counts in cells) and do not address the cosmic error with respect to two-point or higher order statistics such as correlation functions. We defer treatment of these issues to future works.

For a population with a known two-point correlation function $\xi(r)$, it is straightforward to calculate the cosmic variance as a function of cell radius R or equivalently, cell volume V (see e.g. Peebles 1980, or Section 3 below). There are, however, several potential practical difficulties with this simple approach. While the correlation function of galaxies is typically well-approximated by a power law in the strongly non-linear regime ($r \lesssim 10$ -15 Mpc), on larger scales, in the linear regime, the correlation function is ex-

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pected to deviate from the power-law slope measured on smaller scales. Also, estimating the correlation function (especially its slope) of an observed population is more difficult than estimating the number density, so often the latter quantity is known while the former is not. In this situation, we can use the theory of clustering and bias in the Cold Dark Matter (CDM) paradigm to estimate the cosmic variance for a population with a known mean redshift and average comoving number density.

In this *Letter*, we estimate the uncertainty due to cosmic variance for several populations that have been identified in the GOODS survey, and present general results based on CDM theory that can be used to estimate the cosmic variance for populations at $z < 6$. Throughout, we assume cosmological parameters consistent with the recent analysis of *WMAP* data (Spergel et al. 2003): matter density $\Omega_m = 0.3$, baryon density $\Omega_b = 0.044$, cosmological constant $\Omega_\Lambda = 0.70$, Hubble parameter $H_0 = 70$ km/s/Mpc, fluctuation amplitude $\sigma_8 = 0.9$, and a scale-free primordial power spectrum $n_s = 1$.

2. THE GEOMETRY OF GOODS

The Great Observatories Origins Deep Survey (GOODS) covers two fields, the Chandra Deep Field South (CDFS) and the Hubble Deep Field North (HDFN). The CDFS field has dimensions $10' \times 16'$, and the GOODS HDFN field has similar dimensions. For more details and a general overview of the GOODS program, see Giavalisco et al. (2003). Here, we treat the case of a single CDFS-sized field. For widely separated fields, the cosmic variance goes as $1/N_{\text{field}}$, so the variance will decrease by a factor of two when the second field is included. In Fig. 1, we show the comoving volume per unit redshift for several recent, ongoing, and planned deep HST surveys: the original Hubble Deep Field North (Williams et al. 1996), the GOODS CDFS field, the GEMS (Galaxy Evolution from Morphology and SEDs; Rix et al. in prep) field⁴, and the planned Ultra Deep Field (UDF)⁵. For redshifts $z \gtrsim 1$ and $\Delta z \sim 0.5-1$, GOODS samples a volume of a few $\times 10^5$ Mpc³. Fig. 1 also shows the average transverse size ($L = \sqrt{10' \times 16'} = 12.7'$) of the GOODS field as a function of redshift, again compared with the original HDF ($L = \sqrt{5.7 \text{ arcmin}^2} = 2.4'$).

3. THE POWER-LAW MODEL

The relative cosmic variance for a population with known two-point correlation function $\xi(r)$ is given by:

$$\sigma_v^2 = \frac{1}{V^2} \int_0^R \xi(|\mathbf{r}_1 - \mathbf{r}_2|) dV_1 dV_2 \quad (2)$$

(see, e.g. Peebles 1980, p. 234). If the correlation function can be represented by a power-law $\xi(r) = (r_0/r)^\gamma$, then this expression can be evaluated in closed form:

$$\sigma_v^2 = J_2 (r_0/r)^\gamma \quad (3)$$

where $J_2 = 72.0 / [(3 - \gamma)(4 - \gamma)(6 - \gamma)2^\gamma]$ (Peebles 1980, p. 230). Assuming spherical cells, the variance may be equivalently expressed in terms of the cell radius R or the cell volume $V \equiv 4\pi R^3/3$.

For objects with a known correlation function that is well represented by a power-law, we can simply use Eqn. 3 to compute the cosmic variance for a given effective volume, as illustrated in Fig. 2. We show σ_v as a function of volume, for three populations with correlation function estimates from the literature: Extremely Red Objects (EROs) at mean redshift $\bar{z} \sim 1.2$, U-band dropouts at $\bar{z} \sim 3$ (also known as Lyman break galaxies (LBG)), and B-band dropouts at $\bar{z} \sim 4$. The magnitude limit and color selection used for each of these populations selects objects in a given redshift range, resulting in an effective volume V_{eff} . Characteristic number densities for each of these populations, along with correlation function parameters and the relevant references, are summarized in Table 1. For example, for EROs in the GOODS field, $\sigma_v \sim 0.4 - 0.6$, while for the less clustered LBGs, $\sigma_v \sim 0.15 - 0.2$. Note that we have assumed here a spherical geometry for the cells, while in fact for the GOODS survey the cells are very elongated, with the redshift dimension being much longer (about a factor of ten) than the transverse dimension in comoving distance units. We have also ignored the evolution in clustering that occurs over the time interval between the ‘back’ and the ‘front’ of the cell. It should be noted that, for two fields with the same volume, the cosmic variance is *smaller* for an elongated (parallelepiped or cylindrical) field than for a compact (cubical or spherical) field (see e.g. Newman & Davis 2002). This is because an elongated field samples more independent (uncorrelated) regions. Therefore, the estimates given here provide an upper bound on the cosmic variance.

⁴ <http://www.mpia.de/homes/barden/gems/gems.htm>

⁵ <http://www.stsci.edu/science/udf/>

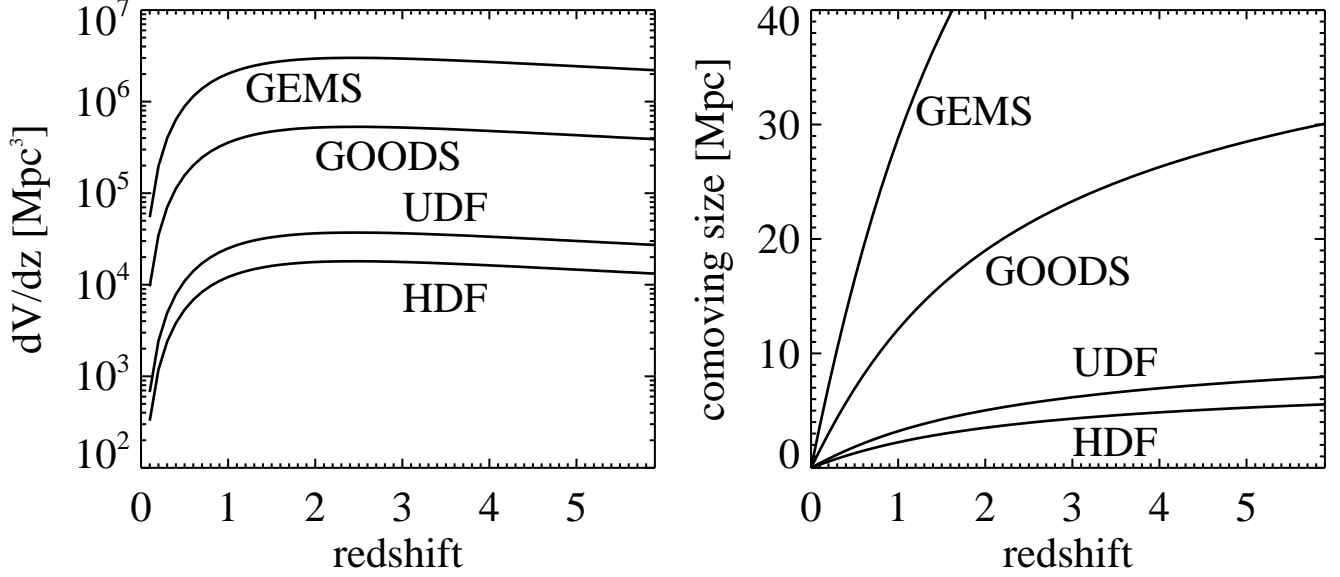


FIG. 1.— [left] The comoving volume per unit redshift spanned by (from bottom to top), the original HDF, the UDF, the GOODS field and the GEMS field. [right] The transverse comoving size of the HDF, UDF, GOODS, and GEMS fields.

TABLE 1
SUMMARY OF PARAMETERS FOR REPRESENTATIVE POPULATIONS

object	\bar{z}	mag. limit	n (h ³ Mpc ⁻³)	r_0 (h ⁻¹ Mpc)	γ	Ref.
ERO	1.2	$K_s < 19.2$	$\sim 10^{-3}$	12 ± 3	[1.8]	D2001
ERO	1.2	$18 < H < 20.5$	$(1.0 \pm 0.1) \times 10^{-3}$	9.5 ± 5	[1.8]	M2001
U-drop	3	$\mathcal{R} < 25.5$	4.7×10^{-3}	3.96 ± 0.29	1.55 ± 0.15	A2003
B-drop	4	$i' < 26$	1.78×10^{-3}	$2.7^{+0.5}_{-0.6}$	[1.8]	O2001

Note. — The mean redshift, magnitude limit, number density, correlation length, and correlation function slope for the populations shown in Fig. 2. References are as follows: D2001 – Daddi et al. (2001); M2001 – McCarthy et al. (2001); A2003 – Adelberger et al. (2003); O2001 – Ouchi et al. (2001). Where the correlation function slope is in brackets, this indicates that the value was assumed in, rather than derived from, the analysis.

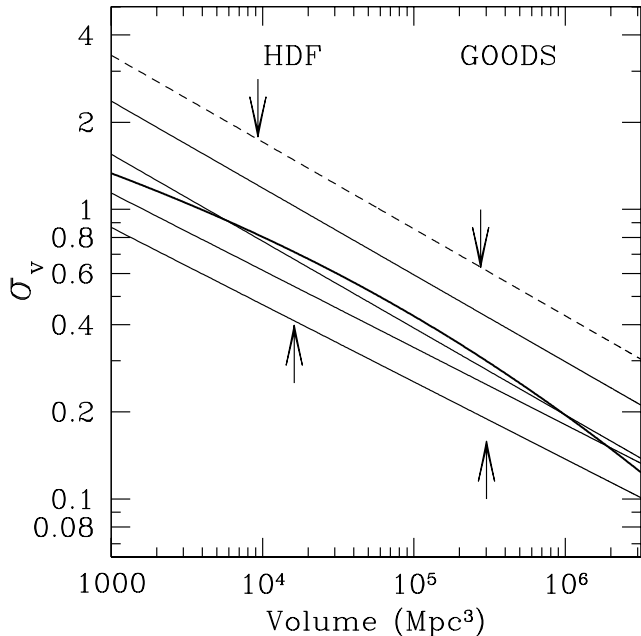


FIG. 2.— The square root of the cosmic variance, σ_v , is plotted as a function of cell volume V . The middle set of lines are for objects which are as clustered as the dark matter at $z = 0$. The slightly curved line shows σ_{DM} from linear theory, while the straight line is a power law model with $r_0 = 5h^{-1}$ Mpc and $\gamma = 1.8$. The topmost solid and dashed lines are for objects as clustered as EROs at $z = 1.2$ (with $r_0 = 9.0h^{-1}$ Mpc and $r_0 = 12h^{-1}$ Mpc, respectively). The second to the bottom line is for objects that cluster like U-dropouts (LBGs) at $z = 3$, and the bottom-most line is for objects as clustered as B-dropouts at $z \sim 4$. The arrows show representative effective volumes for the original HDF and GOODS, for EROs (top set) and LBGs (bottom set).

4. CDM MODELS

We now consider the situation in which we know only the number density but not the correlation function of a population. In this situation, we can use predictions from CDM to estimate the clustering strength of a population with a given number density at a known average redshift. In CDM models, both the number density and clustering strength of dark matter halos are a strong function of halo mass. Fig. 3 shows the average bias as a function of number density, for dark matter halos at various redshifts (as described in the figure caption), computed using the analytic model of Sheth & Tormen (1999). The bias is defined as the ratio of the root variance of the halos and the dark matter, $b \equiv \sigma_h/\sigma_{DM}$. It is likely that this relationship is more complicated for galaxies, since there is probably not a one-to-one correspondence between galaxies and dark matter halos. Similar relations for more general ‘occupation functions’ (i.e., allowing varying numbers of galaxies per halo) are given in e.g. Moustakas & Somerville (2002). Fig. 3 also shows the variance of dark matter σ_{DM} as a function of cell volume for the same redshifts, as predicted by linear theory ($\sigma(R, z) = \sigma(R, 0) D_{lin}$, where D_{lin} is the linear growth function). We now have all the ingredients necessary to obtain a rough estimate of the cosmic variance for any population associated with dark matter halos with a known number density and mean redshift:

1. Read off the average bias b for objects of a given number density and mean redshift from the left panel of Fig. 3.
2. Obtain the value of σ_{DM} at the relevant scale V and redshift from the right panel of Fig. 3.
3. The cosmic variance for the population is then given by $\sigma_v = b \sigma_{DM}$.

As a consistency check, we can use the values given in Table 1 to estimate the cosmic variance for the same populations discussed in the previous section. For EROs, using the number density $n = 1.0 \times 10^{-3} h^3 \text{Mpc}^{-3}$, we would estimate a bias of $b \sim 1.8$, resulting in $\sigma_v \sim 0.7$ at $z = 1$, in reasonable agreement with our earlier estimate of $\sigma_v \sim 0.6$ at $z = 1.2$. Similarly, for LBGs at $z = 3$, we find $b \sim 2.5$, resulting in $\sigma_v \sim 0.25$, again in agreement with the earlier estimate of $\sigma_v \sim 0.2$. One reason that these estimates are not in precise agreement with the values obtained from the calculation based on the actual correlation length is due to the unknown halo occupation distribution (i.e., the number of galaxies per halo as a function of halo mass). It has been shown previously that one cannot simultaneously exactly reproduce both the number density and observed correlation length of either of these populations under the simple assumption of one galaxy per halo (Moustakas & Somerville 2002), adopted here.

5. CONCLUSIONS

Cosmic variance can be a significant source of uncertainty in estimates of the number density or related quantities in deep surveys. We have given empirical estimates of the uncertainty due to cosmic variance for several populations that have been identified in the GOODS survey: EROs at $z \sim 1$, U-dropouts at $z \sim 3$ and B-dropouts at $z \sim 4$. These empirical estimates were based on correlation function measurements from the literature for similarly defined populations, and may be refined once correlation function estimates have been obtained for the actual populations identified in GOODS. From this calculation, we saw that the cosmic variance is much reduced in GOODS compared with the original HDF (40–60% rather than a factor of 2 for for very strongly clustered populations such as EROs, 15–20% rather than 40% for less clustered populations such as LBGs). We have also presented predictions from the theory of clustering and bias in a Λ CDM Universe, which allow one to estimate the cosmic variance for a population of a known average redshift and number density but unknown clustering strength. We emphasize that this approach is intended to give only a simple first order estimate of the cosmic variance. More detailed estimates, tailored to individual populations and including treatments of e.g. a generalized halo occupation distribution formalism, geometric effects, the observational selection function, and clustering evolution and the change in absolute magnitude limit over the redshift range of the sample, will be addressed in future works.

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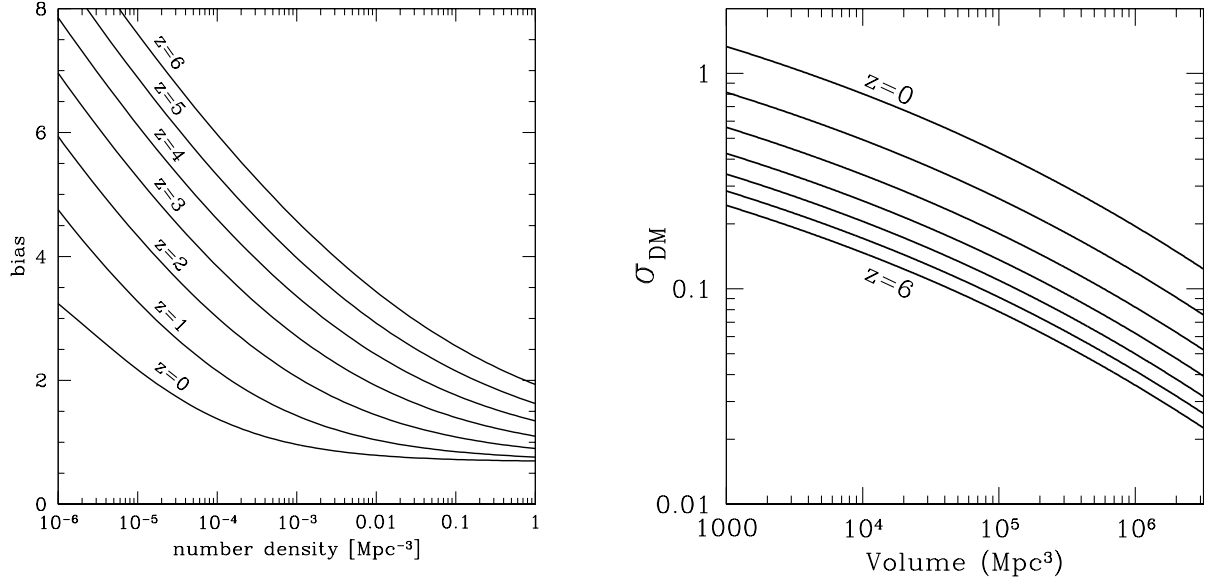


FIG. 3.— [left] Bias as a function of comoving number density, for dark matter halos at $z = 6, 5, 4, 3, 2, 1$, and $z = 0$ (from top to bottom). [right] Variance of dark matter from linear theory, for the same redshifts ($z = 0-6$ from top to bottom).

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REFERENCES

- Adelberger, K. L., Steidel, C. C., Shapley, A. E., & Pettini, M. 2003, *ApJ*, 584, 45
 Colombi, S., Szapudi, I., Jenkins, A., & Colberg, J. 2000, *MNRAS*, 313, 711
 Daddi, E., Broadhurst, T., Zamorani, G., Cimatti, A., Röttgering, H., & Renzini, A. 2001, *A&A*, 376, 825
 Giavalisco, M. et al. 2003, *ApJ*, this volume
 McCarthy, P. J. et al. 2001, *ApJ*, 560, L131
 Moustakas, L. A. & Somerville, R. S. 2002, *ApJ*, 577, 1
 Newman, J. A. & Davis, M. 2002, *ApJ*, 564, 567
 Ouchi, M. et al. 2001, *ApJ*, 558, L83
 Peebles, P. 1980, *The Large-Scale Structure of the Universe* (Princeton University Press)
 Sheth, R. K. & Tormen, G. 1999, *MNRAS*, 308, 119
 Spergel, D. N. et al. 2003, preprint, astro-ph/0302209
 Szapudi, I., Colombi, S., & Bernardeau, F. 1999, *MNRAS*, 310, 428
 Szapudi, I., Colombi, S., Jenkins, A., & Colberg, J. 2000, *MNRAS*, 313, 725
 Williams, R. E. et al. 1996, *AJ*, 112, 1335